

Problem Setting

- Frequently used portfolio allocation processes require the estimation of the covariance matrix of the assets returns (e.g. *Global Minimum Variance* [Maillard 10, Clarke 12], *Maximum Variety* [Fideas Capital] or *Most Diversified Portfolio* [Choueifaty 08], *Mean-Variance* [Markovitz 52], etc.)
 - The Sample Covariance Matrix (SCM) - optimal under the Normal assumption - is the mostly used estimator, but, financial time series of returns might exhibit outliers,
 - The field of robust estimation intends to deal with outliers ([Maronna 76, Tyler 87]),
 - Hybrid robust shrinkage covariance matrix estimates have also been proposed building estimators upon Tyler's robust M-estimator ([Chen 11, Ollila 14, Pascal 14]),
 - Recent works based on Random Matrix Theory (RMT) have also considered robust estimation in the large dimensional regime ([Yang 15]).
- A way to mitigate covariance matrix estimation errors is to identify the most informative asset part and then to filter the noisy part of the data
 - Standard statistical methods like the principal component analysis may fail in distinguishing informative factors from the noisy ones,
 - RMT helps in finding a solution for filtering noise ([Laloux 99 and 00, Potters 05]), but needs to be adapted to non-homogeneous and correlated time series,
 - ⇒ To fill this gap, we extend our preliminary results found in [Jay 18] and based upon the most up-to-date RMT-based model order selection [Vinogradova 13, Terreaux 17] to non-homogeneous time series.

We propose in this paper...

... a new process for estimating and denoising covariance matrix that leads to improved global portfolio performances (reduced Draw-Down, increased Sharpe Ratio, etc.).

- Asset returns are modelled as a multi-factor model ([Jay 11, Darolles 13]),
- Assets are grouped regarding the distribution of their returns,
- The up-to-date RMT-based Model Order Selection method [Terreaux 17, Jay 18] is applied on groups of homogeneous assets to denoise the assets covariance matrix.

Assets returns Model

Let $\{\mathbf{r}_t\}_{t \in [1, N]}$ be N observations of the m assets returns, modelled as a K -factor model. For each observation date t , we then have:

$$\mathbf{r}_t = \sum_{k=1}^K \mathbf{f}_{t,k} \beta_k + \sqrt{\tau_t} \mathbf{C}^{1/2} \mathbf{n}_t, \quad t \in [1, N],$$

or, written more compactly: $\mathbf{R} = \mathbf{B} \mathbf{F} + \mathbf{C}^{1/2} \mathbf{N} \mathbf{T}^{1/2}$, where

- $\mathbf{R} = [\mathbf{r}_1, \mathbf{r}_2, \dots, \mathbf{r}_N] \in \mathbb{R}^{m \times N}$ are the observations,
- $\mathbf{B} = (\beta_1, \dots, \beta_K) \in \mathbb{R}^{m \times K}$ is an unknown mixing matrix of coefficients (or beta) that define the proportion of the K factors in each asset,
- $\mathbf{F} = (\mathbf{f}_1, \dots, \mathbf{f}_N) \in \mathbb{R}^{K \times N}$ is an unknown matrix of the K common returns,
- $\mathbf{T} = \text{diag}(\tau_1, \dots, \tau_N) \in \mathbb{R}^{N \times N}$ is a diagonal matrix unknown containing random texture,
- $\mathbf{N} \in \mathbb{R}^{m \times N}$ is a white Gaussian noise ($E[\mathbf{n}_t^T \mathbf{n}_t] = 1$), independent of the K factors,
- $\mathbf{C} \in \mathbb{R}^{m \times m}$ is an unknown **Toeplitz** scatter matrix ($\text{Tr}(\mathbf{C}) = m$).

Theoretical Results [Terreaux 17]

Robust Consistent Estimation for \mathbf{C}

Let $\hat{\mathbf{M}}_{FP} = \frac{m}{N} \sum_{t=1}^N \frac{\mathbf{r}_t \mathbf{r}_t^T}{\mathbf{r}_t^T \hat{\mathbf{M}}_{FP}^{-1} \mathbf{r}_t}$ be the scatter matrix Tyler M-estimator of \mathbf{R} .
As $m, N \rightarrow \infty$ such that $m/N \rightarrow c \in]0, \infty[$, we have $\|\mathcal{T}[\hat{\mathbf{M}}_{FP}] - \mathbf{C}\| \xrightarrow{a.s.} 0$,
where $\mathcal{T}[\cdot]$ is the **Toeplitz rectification** operator: $(\mathcal{T}[\mathbf{X}])_{ij} = \frac{1}{m} \sum_{k=1}^m \mathbf{X}_{k, k+|i-j|}$.
A consistent estimator $\hat{\mathbf{C}}$ of the background scatter matrix \mathbf{C} characterizing the background noise is therefore defined through observations \mathbf{R} as $\hat{\mathbf{C}} = \mathcal{T}[\hat{\mathbf{M}}_{FP}]$.
⇒ The observations \mathbf{R} can now be whitened through $\hat{\mathbf{C}}^{-1/2} \mathbf{R}$

Behavior of whitened data

Let $\mathbf{R}_w = (\mathcal{T}[\hat{\mathbf{M}}_{FP}])^{-1/2} \mathbf{R}$ be the whitened data and $\hat{\mathbf{W}}_{FP}$ be the Tyler M-estimator of \mathbf{R}_w . As $m, N \rightarrow \infty$ such that $m/N \rightarrow c \in]0, \infty[$, if \mathbf{R}_w does not contain any factor, then:

$$\|\hat{\mathbf{W}}_{FP} - \frac{1}{N} \mathbf{N} \mathbf{N}^T\| \xrightarrow{a.s.} 0.$$

- Without factors, the spectral distribution of the whitened data scatter matrix of \mathbf{R}_w follows a Marchenko-Pastur distribution (same spectral distribution of unobservable covariance matrix of \mathbf{N}) characterized by its support $[(1 - \sqrt{c})^2, (1 + \sqrt{c})^2]$,
- All eigenvalues greater than $\bar{\lambda} = (1 + \sqrt{c})^2$ can be considered as significant factors,
- Given that the eigenvalues are "detectable", then $\hat{\lambda}_{k,N} \xrightarrow{N \rightarrow \infty} \frac{(\lambda_k + \sigma^2)(\lambda_k + \sigma^2 c)}{\lambda_k}$.

Estimation of K the number of factors

Let $(\lambda_k)_{k \in [1, m]}$ be the sorted eigenvalues of $\hat{\mathbf{W}}_{FP}$, then: $\hat{K} = \underset{k}{\text{argmax}} (\lambda_k > \bar{\lambda})$.

Illustration: estimating the correct number of factors

Estimating K is really a challenging problem for many applications where informative signal is embedded in correlated noise. Below, we show how our process allows to detect the $K = 3$ sources embedded in non-Gaussian and strongly correlated noise.

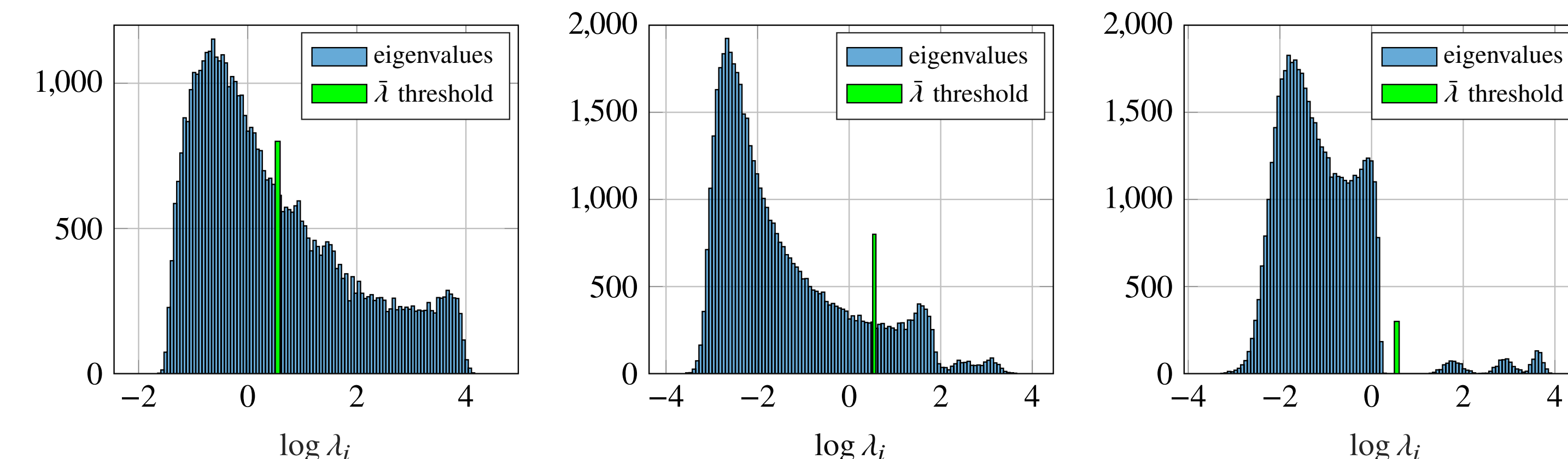


Figure 1: Eigenvalue distributions. Left: $\mathbf{R} \mathbf{R}^T / N$, Sample Covariance Matrix of observations. Middle: $\hat{\mathbf{M}}_{FP}$, Tyler covariance matrix of observations. Right: $\hat{\mathbf{W}}_{FP}$, Tyler covariance matrix of observations after whitening process. K-distributed case with shape parameter $\nu = 0.5$, $\rho = 0.8$, $m = 100$, $N = 1000$, $K = 3$.

Application: the Variety Maximum (VarMax) Portfolio

- VarMax portfolio** weights are the weights that maximise the Diversification Ratio (DR):
 $\mathbf{w}_{vm}^* = \underset{\mathbf{w}}{\text{argmax}} DR(\mathbf{w}) = \underset{\mathbf{w}}{\text{argmax}} \frac{\mathbf{w}^T \mathbf{s}}{(\mathbf{w}^T \boldsymbol{\Sigma} \mathbf{w})^{1/2}}$, where $\boldsymbol{\Sigma}$ is the covariance matrix and \mathbf{s} such that $s_i = \sqrt{\boldsymbol{\Sigma}_{ii}}$, $i \in [1, m]$,
 - The investment universe** consists of $m = 43$ baskets of European equity stocks representing industry subsectors, countries and factor-based indices,
 - Assets are preliminary grouped** into 5 groups regarding the closeness of their respective quantiles,
 - The whitening process** is applied in a "block-manner" for each of the 5 groups.
- We compare three different VarMax portfolios with the MSCI EU Index:
- "RMT-Tyler-Wh-by-Gr": the proposed "by group" whitening process,
 - "RMT-Tyler-Wh": the whitening process on the whole universe,
 - "SCM": the classical SCM estimator is used.

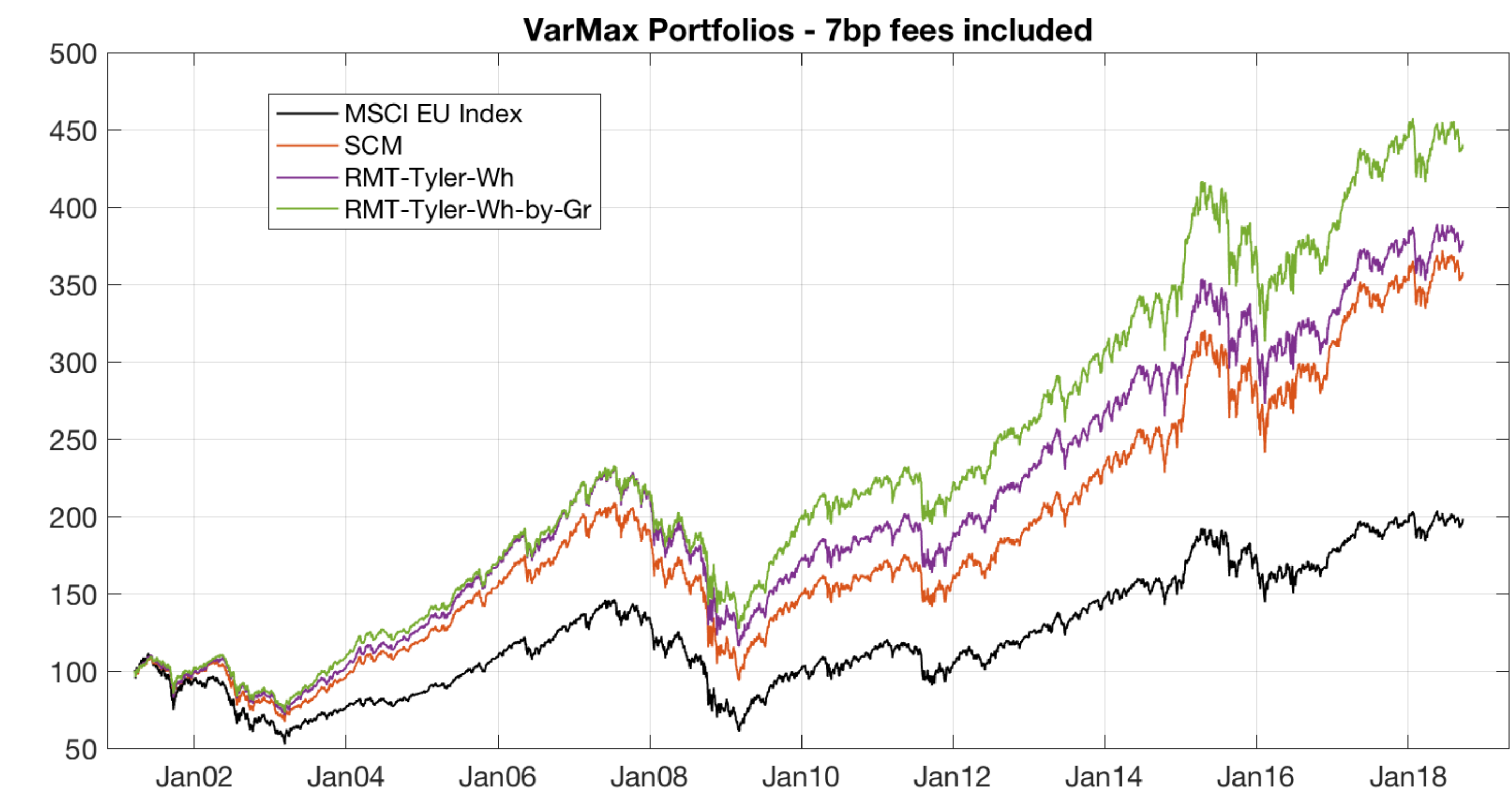


Figure 2: VarMax portfolios wealths from June 2001 to Septembre 2018. The proposed "RMT-Tyler-Wh-by-Gr" (green line) leads to improved performances vs the "RMT-Tyler-Wh" (purple) and the "SCM" (red), as shown in 1: higher annualized return, lower annualized volatility, lower maximum drawdown and higher Diversification Ratio. But it results in a twice higher turnover: we then have taken into account 7bp of transactions fees to compare the portfolio wealths.

Variety Max Portfolios	Annualised Return	Annualised Volatility	Ratio (Return / Volatility)	Maximum Drawdown	Diversification Ratio (avg)
RMT-Tyler-Wh-by-Gr	8.47%	14.58%	0.58	46.27%	1.59
RMT-Tyler-Wh	7.48%	15.20%	0.49	49.22%	1.45
SCM	7.44%	16.17%	0.46	54.70%	1.38
Benchmark	3.41%	19.17%	0.18	58.54%	

Table 1: Some performance numbers for the VarMax Portfolio

Conclusion

- Asset returns have been modelled as a multi-factor model embedded in a correlated elliptical and symmetric noise, allowing to account for non-Gaussian and non correlated noise,
- Given this model setup, the most informative assets have been separated from the noise subspace using a "Toeplitzified" robust and consistent Tyler-M estimator and the Random Matrix Theory applied on the whitened covariance matrix estimate,
- Preliminary grouping assets leads to improved performance, as shown with the Variety Maximum portfolios.